

Using Administrative Record Data to Evaluate the Quality of Survey Estimates

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ABSTRACT

The Survey of Income and Program Participation (SIPP) is a new Census Bureau panel survey designed to provide data on the economic situation of persons and families in the United States. The basic datum of SIPP is monthly income, which is reported for each month of the four-month reference period preceding the interview month. The SIPP Record Check Study uses administrative record data to estimate the quality of SIPP estimates for a variety of income sources and transfer programs. The project uses computerized record matching to identify SIPP sample persons in four states who are on record as having received payments from any of nine state or Federal programs, and then compares survey-reported dates and amounts of payments with official record values. The paper describes the project in detail and presents some early findings.

KEY WORDS: SIPP; Record check; Record linkage; Survey response validity.

1. INTRODUCTION

This paper addresses issues concerning the use of records to evaluate the quality of survey estimates and describes a specific application to the Survey of Income and Program Participation (SIPP) in the United States.

Matching administrative records to survey observations on a case-by-case basis, which we call a “record check,” provides useful information to survey users and designers. A record check enables the analyst to make a full range of measurement error parameter estimates for evaluation purposes. These estimates, in turn, facilitate two basic kinds of activities:

1. quantifying the effects of measurement errors on subject-matter estimates such as means, proportions, correlation coefficients, and multivariate regression coefficients (and possibly adjusting the estimates to correct for the measurement errors), and
2. deriving more efficient survey designs that directly address, for example, the tradeoffs between measurement quality and costs.

1.1 Basic Terms

Our focus here is on measurement (or “response”) errors, although the record check method can be extended to evaluate other nonsampling and sampling errors also. This is not a technical exposition, but we do need to define some of our basic terms first. We assume that the survey observation from sample element i can be expressed as the sum of the true value and an error, e :

$$\text{Survey}_i = \text{True}_i + e_i.$$

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The average bias in a set of N survey observations, which we call the response bias or survey bias, is

$$\bar{e} = \sum e_i/N,$$

and the response error variance is just $\text{Var } e$.

Similarly, the measurement model for the administrative record observation is:

$$\text{Record}_i = \text{True}_i + u_i,$$

so that record bias is \bar{u} and record error variance is $\text{Var } u$.

1.2 Comparison of Evaluation Approaches

The capabilities of the record check approach can be contrasted with other methods of evaluation such as reinterviews and experiments. Reinterviews and other repeated measures designs aim at estimating a very limited set of measurement error parameters, usually something called the simple response variance or the response error variance. These approaches implicitly make strong assumptions about true change over time and about either the true value or bias parameter (Marquis 1986).

One frequently attempted remedy is to create a true value measurement as part of the reinterview program, for example by reconciling discrepant answers with a knowledgeable respondent or by asking much more detailed and specific questions during the reinterview. But the validity of these “true value” measures is suspect. Both Bailar (1968) and Koons (1973) have shown, for example, that reconciled reinterview responses are biased. And while detailed, specific questioning is often preferred to a more global approach, there is no independent evidence that it reduces measurement biases to zero — or at all. Record checks potentially provide higher quality criterion information requiring much weaker (and perhaps more realistic) assumptions for purposes of estimating survey data quality.

A different method of evaluating aspects of surveys is the experiment, such as a fully-crossed factorial design or an interpenetrated design for assigning interviewers. Analysts compare experimental groups with respect to statistics such as subject matter means or proportions and draw conclusions about which treatment produces more or less reporting of the subject matter of interest. What is controversial, however, is determining which is “better” in a measurement sense, a difficulty that is much reduced when criterion data — such as administrative records — are available.

Without criterion data, it is often necessary for the analyst to resort to strong assumptions about measurement errors, such as:

1. more reporting is better reporting;
2. forgetting of meaningful material increases with the passage of time;
3. unbounded interviews contain overreports, bounded interviews don't;
4. reporting performance decays with length of interview or time-in-sample;
5. people are basically lazy and devious — they will lie to avoid being asked a detailed set of questions; and
6. self reports are better than proxy reports.

Indeed, these assumptions have become part of the folklore of survey design in the western world. And yet, it is difficult to find any support for any of these assumptions from appropriately designed record checks. Experiments and related arrangements are excellent approaches to pinpointing the sources of variation, and in untangling estimation problems of collinearity, but are often unnecessary and seldom sufficient for evaluating an existing measurement process.

In sum, these other evaluation approaches are forced to make strong assumptions about:

1. the independence of the original and evaluation measures when they are clearly dependent;
2. the relationship of the original measure to a criterion when no objective, external link exists; and/or
3. cognitive processes not supported by research.

Record checks also employ assumptions in evaluating measurements. For example, the usual way of estimating the response bias is to assume no record bias ($\bar{u} = 0$) and simply calculate the average of the differences between the matched survey and record observed values:

$$\text{Estimated Survey Bias} = \sum (S_i - R_i)/N.$$

While one cannot directly support the no-record-bias assumption, one can conduct meaningful sensitivity tests of the effects of possible violations of the assumption on evaluation conclusions.

1.3 Issues in Designing Record Checks

Several issues merit consideration in designing a record check to evaluate survey measurement. We comment on some of the main ones here: incomplete observation designs, matching errors, record errors, true value differences, and absence of repeated measures or experimental design features.

1.3.1 Incomplete Observation Designs

Past record checks have often used one-directional or partial designs for data collection, such as when we survey people about owning library cards and check the records for those who claim to have one, or sample from a list of people with a diagnosed chronic disease and survey them to see if they report it in a survey questionnaire. Because these partial designs do not observe the full range of response errors in the correct proportions, they yield biased estimates of such classical measurement error parameters as the response bias and the response error variance. One-directional designs can fail to detect some or all of the true survey bias, can cause the analyst to interpret up to one-half of the response error variance as response bias, and can predetermine the sign of the estimated response bias if the measured variable is binary (Marquis 1978). Full designs are a necessary (albeit not sufficient) condition for obtaining unbiased estimates of the desired response errors.

1.3.2 Matching Errors

The essence of the record check is a one-to-one matching of survey and record observations. This is difficult to do correctly, and matching errors (false matches, false nonmatches) will potentially bias the measurement error estimates of interest. Neter *et al.* (1965) show that when there are no unmatched cases, the mismatches will bias the estimates of response error variance upward. In terms of the reliability of a dichotomous measure (which is a function of the response error variance), the estimate will be attenuated by exactly the match error rate (Marquis *et al.* 1986). It is therefore desirable to keep match errors to a minimum and to know something about the errors that remain.

1.3.3 Administrative Record Errors

As noted earlier, one usually has confidence that the records in a record check study are very good measures of the trait of interest. If the implied assumptions about record measurement bias and record measurement error variance are violated, this can cause the response error

estimates to be biased away from zero. For example, bias in the record observations can appear as bias in the survey observations but with the opposite sign. Feather (1972) describes this effect in a record check of physician visits in Saskatchewan, in which an apparently large survey over-reporting rate was due to the record's recording a complete treatment procedure rather than the individual visits for diagnosis. Similarly, the presence of measurement error variance in the record can cause inflated estimates of response error variance in the survey (Marquis 1978).

1.3.4 True Value Differences

Problems arise when the survey and record systems use different definitions. This is often the case in "aggregate comparisons" of population parameter estimates made separately by each source. A common difference is in the scope of the populations covered, such as when the survey frame is limited to the civilian, noninstitutionalized population and the record includes everybody. Case-by-case matching can minimize the threats posed by differential coverage, but even estimates derived from these studies can still be plagued by differences in the concepts or the attributes of the concept. For example, Cox and Iachan (1987) report the results of a study which compared survey-reported health conditions with medical records. The authors conclude that a major reason for the lack of correspondence between survey and record reports was differing concepts — the survey was designed to elicit the complaints which led to doctor visits while the medical records focused on final diagnoses. As an example from our study, the administrative records often contain the date a check was written for a transfer payment, while our survey respondents tell us when they received the payment. Such differences can threaten our time-related estimates of such things as telescoping response errors.

1.3.5 Absence of Experiments and Reinterviews

Evaluation record checks can detect errors but are not good at evaluating the remedies for the errors. To know how well a different survey design might perform, one must usually either test the alternative design options or arrange to estimate parameters of an underlying model from which survey designs can be derived (*e.g.*, a model of forgetting effects). For example, an evaluation record check design can estimate and compare response errors for self and proxy respondents. Without heroic assumptions it cannot, however, suggest how the measurement error parameters would change if the survey's respondent rule were changed (say, to allow only self response).

Similarly, a record check without a reinterview or another set of independent measures is limited in the number of basic error parameters it can estimate. For example, our initial definitions mentioned three parameters: true value, survey error, and record error. Without a reinterview (or other independent measure) there are only two measures with which to estimate the three unknowns. An additional measure can help identify the estimates of the parameters in the model.

2. CHARACTERISTICS OF SIPP

Here we briefly describe the main features of SIPP — the Survey of Income and Program Participation — as a prelude to discussing the record check evaluation design.

2.1 Overview of SIPP Contents

The purpose of SIPP is to provide improved information on the economic situation of people and households in the United States. It collects comprehensive longitudinal data on cash and noncash income, eligibility for and participation in Government transfer programs, assets and

liabilities, labor force participation, and a host of related topics. SIPP data assist the evaluation of the cost and effectiveness of current Federal government programs, the potential impacts of proposed program changes, and the actual impacts of changes when implemented. In general, the Census Bureau and other Government agencies which have fostered and supported the development of SIPP expect it to be an invaluable tool for domestic policy planning (Nelson *et al.* 1985).

Core SIPP questions — repeated in each wave of interviewing — cover labor force participation and amounts and types of income received, including transfer payments and noncash benefits from various programs for each month of the reference period. The core questions cover nearly 50 sources of income, including Government transfer payments from retirement, disability and unemployment benefits, and welfare programs such as Aid to Families with Dependent Children. Information is also gathered on noncash programs such as food stamps, Medicare and Medicaid; private transfers such as pensions from employers, alimony, and child support; ownership of assets that produce income, such as interest, dividends, rent and royalties; and on miscellaneous sources of income, such as estates.

2.2 SIPP Data Collection Design

SIPP started in October 1983 with a sample of approximately 25,000 designated housing units (the “1984 Panel”) selected to represent the noninstitutional population of the United States. In February 1985 a new and slightly smaller panel was introduced. Additional panels are to be introduced each February throughout the life of the survey. Due to budget reductions, the sample size for new panels is currently about 15,000 households.

Each sample household is interviewed by personal visit once every four months for 2-1/2 years, resulting in a total of eight interviews. The reference period for each interview is the four months preceding the interview month. At each visit to the household, each person fifteen years of age or older is asked to provide information about himself/herself. Proxy reporting is permitted for household members not available at the time of the visit. Information concerning proxy response situations is recorded and is available for analytical purposes.

To facilitate field operations, each sample panel is divided into four subsamples (“rotation groups”) of approximately equal size, one of which is interviewed each month. Thus, one “wave” or cycle of interviewing is conducted over a period of four months for each panel. This design produces steady field and processing workloads, but it also means that each rotation group uses a slightly different four-month reference period.

Beginning with the second wave of interviewing in the 1984 panel, SIPP conducts reinterviews with a small sample of households about a subset of items (including program participation). These data are used to check for interviewer falsifications and perhaps to estimate response inconsistencies.

3. RECORD CHECK DESIGN

The purpose of the record check is to provide an evaluation of some of the income data gathered in SIPP. We highlight important features of the design of the record check next, covering the samples, the administrative records, the matching approach, and the analysis.

3.1 Record Check Samples

The SIPP Record Check uses a “full” rather than a one-directional design; that is, the records allow us to validate all observed values in the survey. Design options we did not choose include:

1. checking records only for people who claimed to be participating in a program, or
2. drawing a sample of known recipients and interviewing them to determine how truthfully they report.

Both of these designs are incomplete and will result in biased estimates of the response error parameters.

The Record Check Study restricts attention to a subset of available SIPP data from the 1984 Panel. First, the sample of people is restricted to households in four target states: Florida, New York, Pennsylvania, and Wisconsin. In the 1984 Panel this translates to approximately 5,000 households. Second, the study's sample of time periods includes only the first two waves of the 1984 Panel. Figure 1 illustrates the wave, rotation group, interview month, and reference period structure for the target survey data.

Third, the SIPP Record Check Study focuses on the quality of reciprocity and amount reporting for selected Government transfer programs. It compares survey reports and administrative records for five Federally-administered programs (Federal Civil Service Retirement, Pell Grants, Social Security (OASDI), Supplemental Security Income (SSI), and Veterans' Compensation and Pensions), and four state-administered programs (Aid to Families with Dependent Children (AFDC), food stamps, unemployment compensation, and worker's compensation).

We limited the study to four states — Florida, New York, Pennsylvania, and Wisconsin — in order to keep the study to manageable proportions. Major criteria used to select these states were:

1. the presence of a computerized, accessible, and complete record system for all target programs;
2. a large SIPP sample;
3. reasonable geographic diversity; and
4. a willingness to share individual-level data for purposes of this research.

Thus, the states were selected purposively; no attempt was made to sample states to be representative of the Nation.

We requested from each participating state agency identifying and receipt information for all persons who received income from the target program at any time from May 1983 through June 1984. The identical request was made of the participating Federal agencies, with the exception that only recipients residing in one of the four selected states were to be included in the data extract.

Wave	Rotation Group	Interview Month	Reference Period Months											
			Jun	Jul	Aug	Sep	Oct	Nov	Dec	Jan	Feb	Mar	Apr	
1	1	Oct 83	X	X	X	X								
	2	Nov 83		X	X	X	X							
	3	Dec 83			X	X	X	X						
	4	Jan 84				X	X	X	X					
2	1	Feb 84					X	X	X	X				
	2	Mar 84						X	X	X	X			
	3	Apr 84							X	X	X	X		
	4 ¹	May 84								X	X	X	X	

Figure 1. Survey Structure for Data Included in the SIPP Record Check Study.

¹ Technically, rotation group 4 of the 1984 SIPP Panel was not administered a Wave 2 interview. The "missing" interview was transparent to respondents, however, who were simply given their Wave 3 interview at the time they would have received the Wave 2 interview. For present purposes, the Wave 3 interview for rotation group 4 is identical to the Wave 2 interview for all other rotation groups, and is included in the Record Check Study in order to have two interviews from all sample cases. All references in the text of this paper to "Wave 2" include the Wave 3 interview for this portion of the panel.

As noted earlier, errors in the records can cause problems for record check evaluation studies. Although several of the administrative record files obtained for this project contain very minor deficiencies, only two appear likely to pose major analytical problems: the New York worker's compensation file, and the Veterans' Compensation and Pensions file. Each is known to be incomplete in its coverage of recipients. The New York file excludes an unknown number of cases which were "closed" (*i.e.*, cases which had already been adjudicated and for which payments by a private insurance carrier had already begun) at the time the data base was created several years ago. The Veteran's file excludes the approximately one percent of all recipients whose benefits were sent to a financial or other institution. There are no known coverage problems with any other files.

An unavoidable problem which afflicts all of the administrative files to some extent is the discrepancy between payout date and receipt of payment; obviously, the SIPP respondent reports the latter and has no knowledge of the former, and the reverse is true for the program records. Where the payout date is close to the end of a month it may be difficult to distinguish a forward telescoping error from a legitimate difference between month of payment and month of receipt. Where there are definitional discrepancies, such as this payment date issue, our analyses will attempt to model them explicitly.

4. MATCHING

4.1 Introduction

The quality of matching has important effects on some of the most critical response error estimates, such as the response error variance. Ideally, variables used to match survey and record observations are measured without error and are able to identify an individual uniquely. The ideal, of course, is never realized.

However, the variables we have available to match surveys and records should go a long way toward minimizing the match errors. Some, such as social security number (SSN), uniquely identify an individual even if other information such as address is outdated, garbled, or obliterated or missing. For purposes not directly related to this study (although certainly of benefit to it), the Census Bureau has taken special measures to ensure that SSN information as reported to the SIPP is complete and valid. For all Wave 1 and 2 sample persons, reported SSN's and reports of not having an SSN were verified and, if necessary, corrected, by the Social Security Administration. Sater (1986) estimates that as a result of this operation the SIPP file contains a valid SSN for about 95 percent of SIPP sample persons who have one.

The wealth of other data — last name, first name, house number, street name, apartment designation, city, zip code, sex, and date of birth — is sufficient for high quality matching even in the absence of a unique identifier such as SSN. In addition, to aid us in evaluating the impact of any remaining match errors, the Census Bureau's matcher produces an ordinal measure of the goodness of the match/nonmatch of each survey observation to its appropriate administrative record counterpart.

4.2 The Census Bureau's Computerized Match Procedures

The Record Check Study uses computerized matching procedures applying the theoretical record linkage work of Fellegi and Sunter (1969). The process involves multiple discrete steps, but basically there are four:

1. standardizing the common data fields in the two files which the matcher will examine to determine whether a pair of records is a match or not;
2. sorting the two files into small subsets of records (or "blocks") which constitute a feasible number of pairs to be examined by the matcher;

3. determining and quantifying the usefulness of each data field to be considered in the match for identifying true matched pairs; and
4. implementing the computer algorithms which perform the actual record matching.

4.2.1 Standardization

The Record Check Study processes all data files — both the SIPP files and the administrative record files — through an address standardizer which standardizes the format of various components of an address (*e.g.*, street name, type, and direction; city name; state abbreviation; *etc.*) and parses each component into a fixed data field. Several programs have been developed for this purpose; we use the ZIPSTAN standardizer developed at the Census Bureau.

In addition to the standardization procedures which apply to all data files, many files require modifications to individual data fields to ensure a common format across files for matching. Common examples of variables which pose problems of this type are sex (which can be represented by either an alpha (“m” or “f”) or a numeric (“1” or “2”) code); date of birth (which has many variants — *e.g.*, “mm-dd-yy,” or “cc-yy-mm-dd,” or the Julian format); and name (which may be a single field or which may have separate fields for each component). We prepare custom-made programs for this type of standardization.

4.2.2 Blocking

Blocking — establishing subsets of records for the matcher to examine in searching for matched pairs of records (*e.g.*, Jaro 1985) — is necessary when matching files with large numbers of records. Obviously, the probability of finding all true matches would be highest if, for each record on one file, the entire other file were searched for a match. However, for large files such unrestricted searches for matched records are simply not feasible. Blocking each file into subsets of records makes matching large files feasible, but at the cost of excluding some records from the search; it thus increases the likelihood that some true matches will be missed. Ideal blocking components, therefore, have sufficient variation to ensure the partitioning of the files into many (and therefore smaller) blocks, and are effective match discriminators — that is, nearly always agree in true match record pairs and nearly always disagree in true nonmatch record pairs.

The study uses multiple independent blocking strategies for each pair of files to be matched, thus minimizing the likelihood that a true match pair will escape detection as a result of blocking. One primary blocking strategy employs the first three digits of the United States Postal Service’s five-digit ZIP code and a four-character SOUNDEX code derived from the sample person’s/recipient’s last name. The ZIP code is a sub-state geographic indicator which generally is recorded quite accurately according to Census Bureau matching experts. The SOUNDEX algorithm is widely-used for creating a standard length, standard format code from input character strings of varying lengths; its advantage for blocking purposes is that it minimizes blocking errors due to misspellings, although it cannot eliminate such errors entirely. The second primary blocking arrangement uses the last four digits of the SSN.

4.2.3 Data Field Match Weights

With some variation, the data fields used in the matching of the SIPP and administrative record files include house number, street name, apartment number, city, ZIP code, SSN, sex, date of birth, last name, and first name. Intuitively, these fields are not equally useful in determining whether a particular pair is a match or not — as an obvious example, agreement on sex is not as indicative of a true match as is agreement on SSN. Fellegi and Sunter (1969) include, in their presentation of a general theory of record linkage, discussions of weight calculations

reflecting different data fields' differing discriminating powers and how these weights feed into optimal decision rules. The Census Bureau's Record Linkage Research Staff has developed programs using Newton's method for non-linear systems (see Luenberger 1984) to solve the Fellegi-Sunter equations, and these programs are used in the SIPP Record Check Study to compute final match weights.

4.2.4 The Computer Matcher

The Census Bureau's computer matcher executes the Fellegi-Sunter procedures on a user-defined set of data fields on files sorted (blocked) according to user specifications. For each data field to be considered in the match, the user supplies match weight seed values, defines the type of agree/disagree comparison (whether the fields must be exactly comparable in order for the matcher to treat them as agreeing, or whether only approximate comparability is necessary), and identifies missing value entries and specifies how they are to be treated (included or ignored in the calculation for a composite match weight). The user sets the composite weight cutoff values for matched pairs and nonmatched pairs, and generates the appropriate COBOL program codes to conduct a match through GENLINK, the Census Bureau's Record Linkage Program Generator (LaPlant 1987).

In simple terms, the matcher:

1. searches each data file for comparable blocks of records — that is, records which agree exactly on the designated blocking components;
2. counts the number of records in found blocks to ensure that neither file's block size exceeds the preset maximum;
3. computes a composite match weight for all possible pairs of records in the block;
4. within the block, assigns each record in one file to a paired record in the other file according to a formula which maximizes the total composite weight for all pairs in the block;
5. applies the Fellegi-Sunter decision procedure to determine whether a pair is a match, a nonmatch, or requires further review; and
6. produces a "pointer" file map to the paired records in each file.

5. ANALYSIS

Our goals for the record check study are to estimate selected measurement error parameters for our samples of people, content, and times, and to assess how these errors relate both to each other and to variables that reflect survey design features. Our general plan is to use the matched data to estimate for each dichotomous participation variable:

1. the response bias (using the survey-minus-record difference score);
2. predictors of the response bias (using logistic or probit regression techniques or possibly LISREL techniques based upon matrices containing polyserial and tetrachoric coefficients of association (Jöreskog and Sörbom 1984));
3. the response error variance (*e.g.*, derived from regression residuals);
4. the conditions or groups associated with very large and very small response error variances; and
5. the kinds and amounts of confusion among transfer programs that contribute to the response errors (using covariance structure analysis procedures such as LISREL).

(We will estimate the same parameters for reports of the amounts of money received from each transfer program but have not yet selected our basic estimation approach.)

The time-in-sample and rotation group hypotheses suggest that response errors will be greater in the second interview than the first, after correcting for any seasonal effects. We plan to examine this and, if we find it to be true, test some of the ideas in the literature about why it may be true. Are the sample elements that survive from the first to the second interview different, as Stasny and Feinberg (1985) suggest, or does the quality of the survivors' reporting deteriorate, as the Neter and Waksberg (1966) conditioning hypothesis might predict?

We don't know yet the extent to which SIPP is experiencing these more traditional problems of longitudinal surveys. One problem for which there is evidence, however, concerns the estimation of month-to-month changes in program participation (Burkhead and Coder 1985). Specifically, more changes in program participation take place at the "seam" between interviews (between September and October in Figure 2) than between the months covered by any one interview (*e.g.*, between June and July or July and August or August and September). The Census Bureau has not published monthly program participation transition estimates from SIPP yet because the estimates show a pattern that appears to be affected heavily by measurement error. Moore and Kasprzyk (1984) and Hill (1987) have speculated about what kinds of response, nonresponse, or procedural errors might be producing the pattern and which set of transition estimates is more accurate. By addressing the problem with administrative data, we hope to come much closer to a definitive explanation about the role of response and nonresponse errors in producing the observed pattern.

Related, possibly, to the seam bias issue is the better-understood phenomenon that measurement error variance tends to inflate estimates of gross change or underestimate stability. Recent literature (*e.g.*, Fuller and Tin 1986) suggests several possible approaches to the problem. We plan to begin the empirical exploration of the measurement error effects on the transition estimates to learn whether, for example, we can base corrections for the response errors on estimates from reinterviews.

Finally, we have hinted previously at the problems that may arise in getting unbiased estimates of the errors if the records also contain errors. We plan, with the use of reinterview measures (that identify the estimate of $\text{Var } e$) to estimate the record error variance ($\text{Var } u$). However, we have no plans to relax the assumption that the records are unbiased.

6. PRELIMINARY FINDINGS

To illustrate our approach, we examine the "seam" issue with data for two Government transfer programs in one state. Recall that the seam problem is that monthly survey reports about program participation status show more frequent status changes between months covered by separate interviews than between other months (covered by the same interview). With the administrative record data we are able to begin to answer key questions concerning the quality of SIPP transition estimates: Are too many transitions reported at the seam? Are too few reported for other months? Does SIPP capture the right number of changes over the whole reference period but distribute them incorrectly?

Figures 3 and 4 contain results of our initial seam bias analyses. Data for these initial analyses come from matched/merged SIPP and administrative record files for Aid to Families with Dependent Children (AFDC) and food stamps in the state of Wisconsin.

A total of 1,632 people were eligible SIPP sample persons in Wisconsin in Wave 1 of the 1984 SIPP Panel. Of this total, 92 (6%) refused to report an SSN and were excluded both from the administrative record match and from the response error analyses. Also, the sample residing in Wisconsin is part of a national sample and is not necessarily representative of Wisconsin.

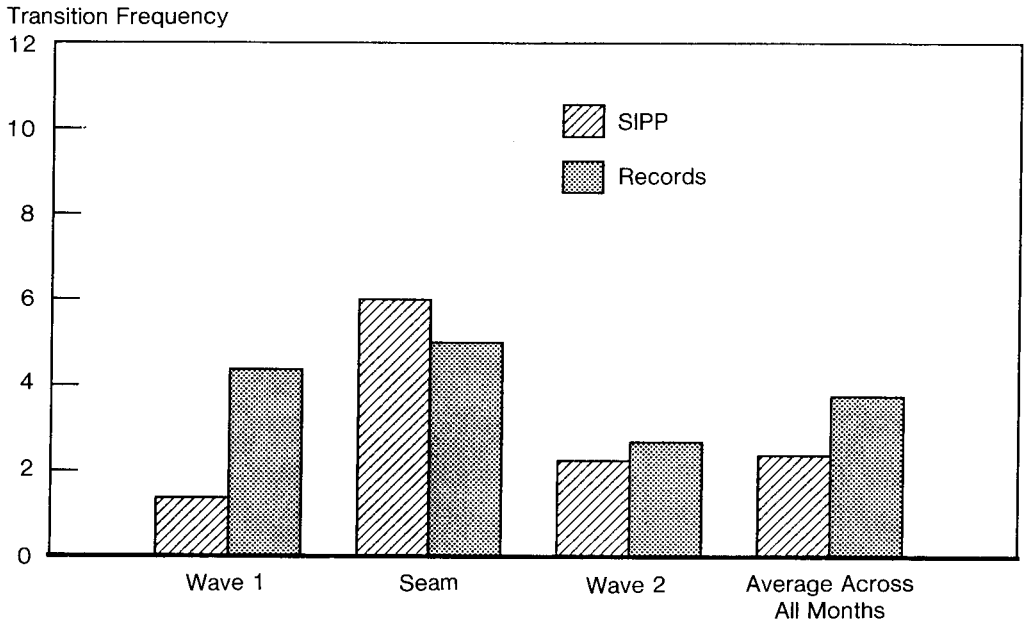


Figure 3. Month-to-Month AFDC Participation Transitions: Comparison of Transition Frequency at the Seam with the Average Frequency Within Waves 1 and 2, and the Overall Average Across All Months, for SIPP and Administrative Records.

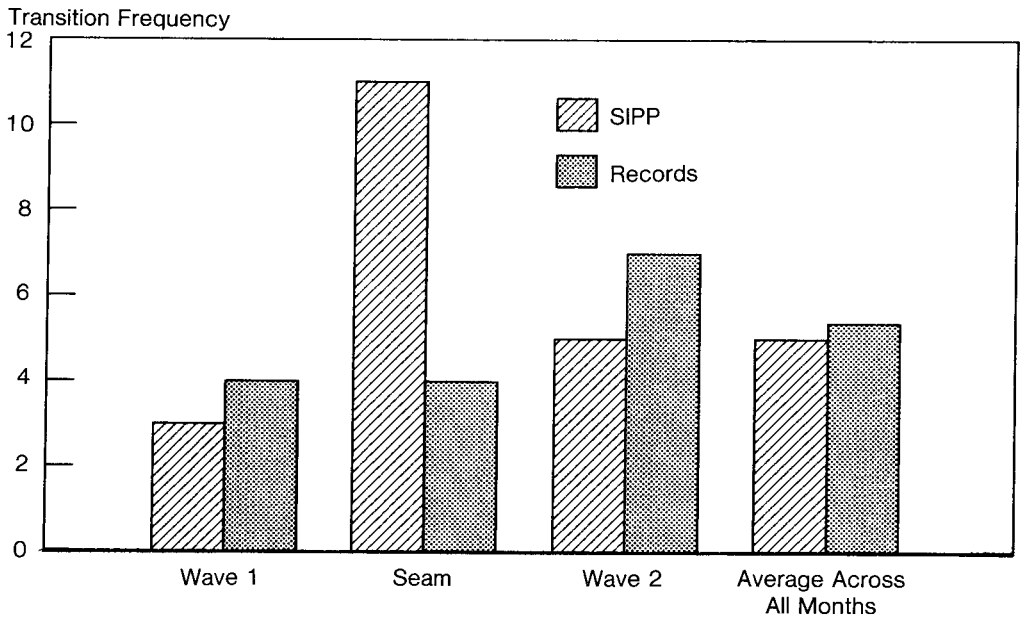


Figure 4. Month-to-Month Food Stamps Participation Transitions: Comparison of Transition Frequency at the Seam with the Average Frequency Within Waves 1 and 2, and the Overall Average Across All Months, for SIPP and Administrative Records.

SIPP procedures assume that all sample persons identified in Wave 1 were eligible sample persons in the same household for all months of the Wave 1 reference period, and that no one other than those eligible at the Wave 1 interview was a household member in the preceding four months. Thus, the month-to-month transition estimates within Wave 1 derive from a constant respondent base of $(1,632-92 =) 1,540$ people. In Wave 2, however, the fluidity of household composition is recognized, resulting in respondent bases which vary slightly from one month-pair to the next — including the interview seam. In the data below the number of eligible persons in both “seam” months is 1,517; within Wave 2 the respondent bases for the three month-pairs are 1,522, 1,531, and 1,532. (Separate analyses (not shown here) indicate that the trends shown in Figures 3 and 4 are not sensitive to excluding people not present in all eight months of the Wave 1 and 2 reference periods.) Because of the small number of cases and the unrepresentative nature of the Wisconsin sample we do not offer inferential statistics for this set of illustrations.

In the figures, the striped bars indicate the number of transitions according to administrative records and the empty bars indicate the number of transitions according to SIPP. If there are too many SIPP transitions at the seam, the empty bar should tower over the striped bar for the comparisons labelled “Seam.” If there are too few transitions reported in SIPP for the months covered within an interview, the empty bar should be smaller than the striped bar for the comparisons labelled “Wave 1” and “Wave 2.” And, if SIPP interviews yield approximately the right number of transition reports, the empty and striped bars should be approximately the same height for the comparisons labelled “Average Across All Months.”

Figure 3 presents the average frequency of month-to-month transitions in Wisconsin AFDC participation within Waves 1 and 2 for the two data sources, and contrasts those figures with the number observed at the Wave 1/2 interview seam. The SIPP “seam bias” problem is quite apparent — the frequency of transitions at the seam is greater than the average within either interview. Although the absolute differences with this sample size are small, the record data suggest that the AFDC seam bias results from a combination of too many transitions reported at the seam and too few in the within-interview months. The final columns of Figure 3 suggest, additionally, a net underreporting of AFDC transitions in SIPP, in addition to the time placement problem.

The Wisconsin food stamps results are summarized in Figure 4, where the seam bias effect in SIPP is even clearer. Once again, the administrative record data suggest a tendency for within-interview transitions to be consistently underestimated with SIPP data. And, in this instance the contrast of survey and record data is even more clear in indicating that SIPP seam transitions are severely overestimated. Unlike the AFDC results, however, both survey and record contain about the same number of transitions overall, suggesting just a time placement problem and not a net underreporting bias.

7. CONCLUSIONS

After a lengthy matching and file preparation process, we are just beginning our analysis of this rich data set. However, with just the initial results presented here we have already shown how record check findings can contribute to our understanding of important measurement error issues — in this case, the SIPP seam bias. There are many more tests to be done and many hypotheses to explore before we can draw definitive conclusions about the nature of SIPP measurement errors and their probable causes. We are confident that the SIPP Record Check Study will allow us to make important advances toward understanding the sizes and forms of these survey errors and perhaps suggest their causes.

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