Component of Statistics Canada Catalogue no. 82-003-X Health Reports



### The feasibility of establishing correction factors to adjust self-reported estimates of obesity

by Sarah Connor Gorber, Margot Shields, Mark S. Tremblay and Ian McDowell





Statistics Statistique Canada Canada Canada



# The feasibility of establishing correction factors to adjust self-reported estimates of obesity

by Sarah Connor Gorber, Margot Shields, Mark S. Tremblay and Ian McDowell

#### Abstract

#### Background

This study examines the feasibility of developing correction factors to adjust self-reported measures of body mass index (BMI) to more closely approximate measured values.

#### Data and methods

Data are from the 2005 Canadian Community Health Survey (subsample 2), in which respondents were asked to report their height and weight, and were subsequently measured. Regression analyses were used to determine which socio-demographic and health characteristics were associated with the discrepancies between self-reported and measured values. The sample was then split into two groups. In the first, self-reported BMI and the predictors of the discrepancies were regressed on measured BMI. Correction equations were generated using all predictor variables that were significant at the p<0.05 level. These correction equations were then tested in the second group to derive estimates of sensitivity, specificity and obesity prevalence. Logistic regression was used to examine relationships between self-reported, measured and corrected BMI and obesity-related health conditions.

#### Results

Corrected estimates provide more accurate measures of obesity prevalence, mean BMI and sensitivity levels (percentage correctly classified). In almost all cases, associations between BMI and health conditions are more accurate when based on corrected versus selfreported values.

#### **Keywords**

Bias, body mass index, direct measure, measurement error, obesity, overweight, prevalence, self-report

#### Authors

Sarah Connor Gorber (613-951-1193; Sarah.ConnorGorber@statcan.gc.ca) and Margot Shields (613-951-4177; Margot.Shields@statcan.gc.ca) are with the Health Information and Research Division at Statistics Canada, Ottawa, Ontario, K1A 0T6. Mark S. Tremblay is with the Children's Hospital of Eastern Ontario Research Institute. Ian McDowell is with the Department of Epidemiology and Community Medicine, University of Ottawa. Obesity is a public health problem in both the developed and developing world. Globally, an estimated 400 million people are obese.<sup>1</sup> In Canada, the prevalence is estimated to be 23% in adults<sup>2</sup> and 8% in children,<sup>3</sup> with rates expected to rise in coming years.<sup>4,5</sup> The costs associated with obesity represent approximately 2% of Canada's total health care expenditures.<sup>6</sup>

Because collecting measured data is expensive, national estimates of the prevalence of obesity are usually based on self-reported survey data. In most countries, body mass index (BMI) is used to estimate the prevalence of obesity because BMI can be easily calculated from self-reported height and weight. However, in both clinical and population samples, self-reports have tended to overestimate height and underestimate weight, which results in a systematic underestimation of obesity prevalence.7-10 This tendency has recently been confirmed in a review of 64 international studies.<sup>11</sup> as well as in Canadian research.<sup>12</sup>

Underestimating the prevalence of obesity is important not only because obesity itself can cause social and physical impairment, but also because it is a risk factor for disease.<sup>13-15</sup> When estimates of obesity are based on self-reported data, the relationship between obesity and conditions such as diabetes, hypercholesterolemia, hypertension, arthritis and heart disease is substantially exaggerated.<sup>16-18</sup>

Given that much population health surveillance will continue to rely on self-reported data, it has been suggested19 that estimates of obesity based on self-reports could be adjusted to more closely approximate measured values. Using data from the 2005 Canadian Community Health Survey (CCHS), which collected both selfreported and measured height and weight, this study examines the feasibility of developing correction equations to adjust self-reported estimates, and assesses whether these equations improve the estimation of obesity (when based on BMI).

#### Methods

#### **Data source**

Data for this study come from the 2005 CCHS. The CCHS is an ongoing survey designed to provide timely crosssectional estimates of health determinants, health status and health system use at a sub-provincial level.<sup>20</sup> Sampling is based on a multi-stage cluster sampling technique that is representative of over 98% of the Canadian population (members of the Canadian Forces, individuals living on Indian Reserves or Crown lands, and residents of institutions, Canadian Forces bases and certain remote regions are excluded). Three sampling frames were used to select households for the 2005 survey: 49% of the sample of households came from an area frame; 50% from a list frame of telephone numbers; and the remaining 1% from a Random Digit Dialing (RDD) sampling frame. More details about the CCHS are available in Béland, 2002.<sup>21</sup>

The 2005 CCHS collected data from 132,947 respondents, yielding a response rate of 79%. A subsample of 7,376 respondents aged 12 years or older were asked their height and weight, and later in the interview, were directly measured. These respondents were all drawn from the area frame and were selected across the ten provinces in proportion to their populations (residents of the territories were excluded). Measured height and weight were obtained for 4,735 individuals—a response rate of 64%. (The main reason for non-response was refusal.)

Because of the high non-response to measured height and weight, an adjustment was made to minimize nonresponse bias. A special sampling weight was created by redistributing the sampling weights of non-respondents to respondents using response propensity classes. The variables used to create these classes were region (British Columbia, Prairies, Ontario, Quebec, Atlantic provinces), age, sex, household size, marital status, rural/urban indicator, and quarter of data collection.

The present study included only adults aged 18 years or older. Children are in a stage of development where weight and height may change over short periods of time. It has also been suggested that the nature of the reporting error in children and adolescents may differ from that in adults.<sup>10</sup> Women who were pregnant (n=47) or breastfeeding (n=58) were also excluded, as BMI is not recommended for use in these groups. Respondents for whom the difference between selfreported and measured estimates of height, weight or BMI were more than 3 standard deviations from the mean were considered outliers and were excluded from the analysis (n=43, n=44 and n=39, respectively). This left 4,080 respondents with self-reported and measured values for height and weight.

CCHS interviewers were trained to measure height and weight. Height was measured to the nearest 0.5 cm (without shoes) with a measuring tape attached to the wall. Weight was measured to the nearest 0.1 kg (without shoes) with a calibrated digital scale (ProFit UC-321 made by Lifesource). The interview lasted approximately 50 minutes and took place in the respondent's home. Self-reported height and weight were collected near the beginning of the interview; the measurements were taken near the end. Respondents were not told that they would be measured.

Self-reported height and weight were collected with the questions: "How tall are you without shoes on?" and "How much do you weigh?" Categories for height in feet and inches were listed on the questionnaire with corresponding metric values in brackets. Interviewers rounded up to the closest inch for respondents who reported half-inch measures. If questioned, interviewers told respondents to report their weight without clothing. Respondents were asked if they had reported in pounds or kilograms; 94% reported in pounds.

#### **Analytical techniques**

The first step was to use the full subsample (n=4,080) to determine which factors were associated with the bias between self-reported and measured height and weight. The bias was calculated by subtracting the measured value from the self-reported value. Negative values indicated underestimation; positive values, overestimation.

Multiple linear regression was used with the bias as the dependent variable in the model. Socio-demographic and health variables, selected based on a review of the literature, were entered as independent variables. Separate models were estimated with the bias in weight, height and BMI as dependent variables. All models were estimated separately for men and women, because the bias differs between the sexes.<sup>8,22-24</sup> Variables that were significant (p<0.05) were used to develop the correction equations.

The sample was then randomly divided into two parts: split-sample A and split-sample B, each containing approximately 50% of the respondents (2,029 or 49.7% and 2,051 or 50.3%, respectively). Split-sample A was used to generate the correction equations using the variables that were significantly associated with the bias in height, weight and BMI identified in the first step. Split-sample B was used to test the equations. To generate the correction equations, the measured value was the dependent variable, and the self-reported value and any variables that were significantly associated with the bias from the first part of the study were independent variables. Only significant independent variables (or categorical variables for which at least one category was significant) were retained for the final correction equations.

Four models were tested: two Full Models and two Reduced Models. In Model 1 (the first Full Model), estimates of height and weight were first adjusted based on the predictors that were significantly related to the bias in height and weight, respectively, in step 1. BMI was then calculated using the adjusted values of height and weight. In Model 2 (the second Full Model), BMI was adjusted by regressing the predictors of the bias in BMI from step 1 directly onto measured BMI. The Reduced Models were similar, except only selfreported height, weight and BMI were used as independent predictors of the measured values. The models are shown in Table 1.

All analyses were run for men and women separately. Interactions and quadratic terms were tested as appropriate. All variables were entered into the models simultaneously, but only significant variables were retained to generate the final correction equations. Final models were tested to ensure they met the assumptions of independence, linearity, equal variance, and normality.

The correction equations generated from split-sample A were applied to the data in split-sample B. Descriptive statistics (means, prevalence of selected categories) were used to compare the self-reported, measured and corrected estimates of obesity. Sensitivity (proportion of obese, overweight or normal weight respondents, based on measured values, who were classified as obese based on self-reported and corrected estimates) and specificity (proportion of non-obese, nonoverweight or non-normal weight respondents who were correctly classified based on self-reported and corrected estimates) were used to determine if the corrected estimates improved BMI classification, compared with self-reported estimates. According to the World Health Organization<sup>25</sup> and Canadian classification guidelines,<sup>26</sup> respondents were categorized as underweight (BMI less than  $18.5 \text{ kg/m}^2$ ), normal weight (BMI 18.5 to 24.9 kg/  $m^2$ ), overweight (BMI 25.0 to 29.9  $kg/m^2$ ) or obese (BMI 30.0  $kg/m^2$  or more).

Logistic regression was then used to determine if the corrected estimates more accurately modeled the relationship

### Table 1 Correction equations to adjust self-reported estimates of weight, height and body mass index (BMI), Full and Reduced Models

	Equation
Full Models	
Model 1 (Height and Weight)	$\label{eq:weight_measured} \begin{split} \text{Weight}_{measured} = b_0 + b_1(\text{weight}_{self\text{-reported}}) + b_2(\text{var1}) + b_3(\text{var2}) + b_j(x_j) \dots + \text{error} \\ \text{Height}_{measured} = b_0 + b_1(\text{height}_{self\text{-reported}}) + b_2(\text{var1}) + b_3(\text{var2}) + b_j(x_j) \dots + \text{error} \end{split}$
Model 2 (BMI)	$BMI_{measured} = b_0 + b_1(bmi_{self\text{-}reported}) + b_2(var1) + b_3(var2) + b_i(x_i) \dots + error$
Reduced Models	
Model 3 (Height and Weight)	$\label{eq:weight_measured} \begin{split} \text{Weight}_{measured} = b_0 + b_1 (\text{weight}_{self\text{-}reported}) + \text{error} \\ \text{Height}_{measured} = b_0 + b_1 (\text{height}_{self\text{-}reported}) + \text{error} \end{split}$
Model 4 (BMI)	$BMI_{measured} = b_0 + b_1(bmi_{self\text{-reported}}) + error$

between obesity and obesity-related health conditions than did the selfreported estimates. All models controlled for age and sex and examined the relationship between BMI (selfreported, measured and corrected) and one of six conditions: diabetes, heart disease, hypertension, arthritis, activity limitations, and fair or poor self-rated health. The analysis was restricted to respondents aged 40 years or older, because the six conditions are more prevalent in that age range.

Data were appropriately weighted, and all measures of variance were estimated with the bootstrap technique to account for the complex survey design.<sup>27-28</sup> SAS (version 9.1) was used for all analyses.

#### Definitions

The socio-demographic variables included *age* (divided into seven groups: 18 to 24, 25 to 34, 35 to 44, 45 to 54, 55 to 64, 65 to 74, and 75 years or older); *level of education* (less than secondary graduation, secondary graduation, some postsecondary, and postsecondary graduation); *geographic region* (Atlantic, Québec, Ontario, West and British Columbia); *urban or rural area*; *employment status* the week before the interview (full-time, part-time or not working); *immigrant status* (10 or fewer years in Canada, more than 10 years in Canada and Canadian-born); *ethnicity* (collapsed because of sample size into White, East and South East Asian, and Other); and *household income*. Household income groups were derived by dividing total household income from all sources in the previous 12 months by Statistics Canada's low-income cutoff (LICO) specific to the number of people in the household, the size of the community, and the survey year. These adjusted income quotients were grouped into deciles.

The health variables were self-reported health status and mental health status (dichotomized into fair/ poor versus good/very good/excellent); activity limitations imposed by a longterm health problem (sometimes/often versus never); smoking status (daily/ occasional versus non-smoker); selfperceived stress (most days are quite a bit/extremely stressful versus a bit/ not very stressful); life satisfaction (dissatisfied/very dissatisfied versus satisfied/very satisfied); perception of weight (overweight, underweight, or about right); number of physician consultations in the past year (continuous); and chronic conditions arthritis/rheumatism, (asthma, hypertension, diabetes, heart disease, cancer, mood disorders). Sample sizes were too small to examine associations with eating disorders.

Leisure-time physical activity level was based on total energy expenditure (EE) during leisure time. EE was calculated from the reported frequency and duration of all of a respondent's leisure-time physical activities in the three months before the 2005 CCHS interview and the metabolic energy demand (MET value) of each activity, which was independently established.<sup>29</sup>

 $EE = \Sigma(N_i * D_i * MET_i / 365 \text{ days}),$ where

N<sub>i</sub> = number of occasions of activity i in a year,

 $D_i$  = average duration in hours of activity i, and

 $MET_i = a$  constant value for the metabolic energy cost of activity i. An EE of 3 or more kilocalories per kilogram per day (KKD) was defined as *active*; 1.5 to 2.9 KKD, *moderately active*; and less than 1.5 KKD, *inactive*.

The influence of *end-digit preference* (the tendency to round responses to numbers ending in 0 and 5) was examined for weight, because past research has associated it with a reporting bias.<sup>8,10,30</sup> The majority of CCHS respondents (73% of men and 67% of women) reported values for their weight that ended in 0 or 5, although it would be expected that, by chance, only about 20% of respondents would have end-digits of 0 or 5.

#### Results

Consistent with past research, mean values of self-reported height were *over*estimated, while weight and BMI were *under*estimated. Men *over*estimated their height by 1.08 cm, and *under*estimated their weight by 1.84 kg and hence, their calculated BMI by 0.94 kg/m<sup>2</sup>. For women, height was *over*estimated by 0.56 cm, and weight and BMI were *under*estimated by 2.47 kg and 1.19 kg/m<sup>2</sup>, respectively.

The regression results derived from split-sample A that were used to establish the correction equations for *weight* are shown in Table 2. In the Full Models for men, self-reported weight, age and the respondents' perception of being over- or underweight were significant predictors of measured weight. Those who perceived themselves as overweight tended to underestimate their weight, and those who perceived themselves as underweight tended to overestimate their weight; the model adjusted these values up or down as appropriate. The adjusted  $R^2$  was .95 for both the Full and Reduced Models.

For women, factors associated with measured weight were self-reported weight, the perception of being overweight, and end-digit preference (the model added a positive adjustment to self-reported weight to compensate for this tendency). The adjusted R<sup>2</sup> for women for both the Full and Reduced Models was .97.

Results for *height* are found in Table 3. Among men, self-reported height, age and life dissatisfaction were significant predictors of measured height, with a negative adjustment related to age and a positive adjustment for those who reported being dissatisfied with their lives. The adjusted R<sup>2</sup> was .82 for the Full Model and .81 for the Reduced Model. For women, all age groups were significantly associated with measured height except for 45 to 54 years. Also significant were those whose ethnicity was a group other than White or East/South East Asian, and those who reported an activity limitation.

For *BMI* (Table 4), the Full Models adjusted self-reported estimates down for men who were dissatisfied with life and who perceived themselves as underweight, and positive adjustments were made for age. For women, significant predictors of measured BMI were self-reported BMI, education, perception of being overweight, and end-digit preference. The R<sup>2</sup> was higher for the female than the male models, but in both cases, was similar for the Full and Reduced Models.

#### Table 2

Regression results for establishing correction equations for weight, by sex, Full and Reduced Models generated from split-sample A, household population aged 18 years or older, 2005

	Variable	Coefficient	95% confidence interval
Man			
Men Full Model R <sup>2</sup> = 0.95 R <sup>2</sup> (adj)= 0.95	Intercept Self-reported weight (kilograms) Aged 25 to 34 years <sup>†</sup> Aged 35 to 44 years <sup>†</sup> Aged 45 to 54 years <sup>†</sup> Aged 65 to 74 years <sup>†</sup> Aged 75 years or older <sup>†</sup> Perceives self as overweight Perceives self as underweight	-0.30 1.01* 0.54 0.39 0.50 1.69* 0.83 0.39 1.16* -1.52*	-2.7 to 2.1 1.0 to 1.0 -0.5 to 1.6 -0.7 to 1.5 -0.5 to 1.5 0.6 to 2.8 -0.2 to 1.8 -0.6 to 1.4 0.4 to 1.9 -2.9 to -0.1
<b>Reduced Model</b> R <sup>2</sup> =.0 95 R <sup>2</sup> (adj)=0 .95	Intercept Self-reported weight (kilograms)	-2.19* 1.05*	-4.3 to -0.1 1.0 to 1.1
Women Full Model R <sup>2</sup> =0 .97 R <sup>2</sup> (adj)= 0.97	Intercept Self-reported weight (kilograms) Perceives self as overweight End-digit preference (0 and 5)	-1.25 1.04* 1.25* 0.52*	-3.3 to 0.7 1.0 to 1.1 0.5 to 2.0 0.0 to 1.0
<b>Reduced Model</b> R <sup>2</sup> = 0.97 R <sup>2</sup> (adj)=0 .97	Intercept Self-reported weight (kilograms)	-2.14* 1.07*	-3.9 to -0.4 1.0 to 1.1

<sup>†</sup> reference group is ages 18 to 24 years

\* p<0.05

Note: Dependent variable is measured weight.

Source: 2005 Canadian Community Health Survey.

To generate the final equations, adjustments were made for all of the variables in Tables 1 to 3. The final equations are shown in Table 5.

These equations were applied to data in split-sample B to generate corrected estimates of mean height, weight and BMI (Table 6). In all cases, self-reported estimates were statistically different from the measured values, and the corrected estimates were closer than the self-reported estimates to the measured values. In all but one case (the difference in BMI for females in Model 3), the corrected and measured means were not statistically different.

Among men, the proportion who were obese was 13.8% according to selfreported data and 23.1% according to measured data (Table 7); the corrected data generated estimates ranging from 19% to 22%. Self-reported, measured and corrected data yielded similar rates of overweight among men. However, self-reported data overestimated the percentage of men in the normal weight range; the corrected data reduced this bias by 9 to 11 percentage points, with the result that the corrected and measured estimates were similar.

Among women, the proportion who were obese was 12.5% according to self-reported data and 18.9% according to measured data; the corrected data generated estimates ranging from 18.2% to 18.7%. Similarly, for overweight, corrected values were closer than selfreported values to the measured

#### Table 3

**Regression results for establishing correction equations for height, by sex,** Full and Reduced Models generated from split-sample A, household population aged 18 years or older, 2005

	Variable	Coefficient	95% confidence interval
Men			
Full Model	Intercept	12.17*	5.6 to 18.8
$R^2 = 0.82$	Self-reported height (centimetres)	0.93*	0.9 to 1.0
$R^{2}(adi) = 0.82$	Aged 25 to 34 years <sup>†</sup>	-1.48*	-2.4 to -0.5
	Aged 35 to 44 years <sup>†</sup>	-0.43	-1.5 to 0.6
	Aged 45 to 54 years <sup>†</sup>	-1.23*	-2.3 to -0.1
	Aged 55 to 64 years <sup>†</sup>	-2.44*	-3.4 to -1.5
	Aged 65 to 74 years <sup>†</sup>	-2.87*	-4.1 to -1.6
	Aged 75 years or older <sup>†</sup>	-2.84*	-4.2 to -1.5
	Dissatisfied with life	2.22*	0.3 to 4.1
Reduced Model	Intercept	7.70*	0.7 to 14.7
R <sup>2</sup> = 0.81 R <sup>2</sup> (adj)= 0.81	Self-reported height (centimetres)	0.95*	0.9 to 1.0
Women			
Full Model	Intercept	14.85*	9.2 to 20.4
R <sup>2</sup> = 0.83	Self-reported height (centimetres)	0.91*	0.9 to 0.9
R <sup>2</sup> (adj) = 0.83	Aged 25 to 34 years <sup>†</sup>	-1.20*	-2.0 to -0.4
	Aged 35 to 44 years <sup>†</sup>	-0.87*	-1.7 to -0.1
	Aged 45 to 54 years <sup>†</sup>	-0.59	-1.6 to 0.4
	Aged 55 to 64 years <sup>†</sup>	-1.34*	-2.6 to -0.1
	Aged 65 to 74 years <sup>†</sup>	-1.42*	-2.4 to -0.5
	Aged 75 years or older <sup>†</sup>	-3.79*	-5.0 to -2.5
	East or South East Asian <sup>‡</sup>	-0.32	-1.8 to 1.2
	Other ethnicity <sup>±</sup>	-0.73*	-1.4 to -0.1
	Activity limitation	-0.66*	-1.3 to 0.0
Reduced Model	Intercept	8.05*	2.5 to 13.6
R <sup>2</sup> = 0.81 R <sup>2</sup> (adj) = 0.81	Self-reported height (centimetres)	0.95*	0.9 to 1.0

reference group is ages 18 to 24 years

reference group is White

n<0.05

Note: Dependent variable is measured height.

Source: 2005 Canadian Community Health Survey.

prevalence, with a slight 1-to 2percentage-point overestimate in the corrected values. Sample sizes in the underweight category were too small to generate reliable estimates.

Sensitivity values in the normal weight category for self-reported data were 93.9% for men and 91.8% for women (Table 8), meaning that in most cases, self-reports correctly classified people of normal weight into the normal weight category.

Sensitivities for the overweight and obese categories fell to 71.1% and 58.7%, for men, and to 62.6% and 68.5% for women. When the data were corrected, sensitivities increased: the corrected numbers accurately classified as many as 86.1% of obese women, 76% of obese men, 79.7% of overweight women, and 82.8% of overweight men. However, the corrected estimates reduced sensitivities for those in the normal weight range.

Specificities were highest for the underweight and obese categories (Table 8), indicating that it is rare for someone to be classified into these groups based on self-reports unless they actually are underweight or obese.

Table 9 displays adjusted odds ratios relating self-reported, measured and corrected BMI to six obesity-related health conditions. An earlier study<sup>16</sup> demonstrated that self-reported BMI exaggerates the relationship between obesity and these health conditions. Unique to the present analysis is that the models have been re-generated based on the corrected estimates. Compared with the odds ratios from the selfreported models, the odds ratios for the corrected models are reduced in most cases (that is, they are closer to the measured values). Arthritis is an exception, with the corrected estimates inflating the relationships for those who are overweight or obese (class II or II - BMI 35 or more kg/m<sup>2</sup>) even more than what they would be if based on self-reports. In addition, the odds ratios for obese class I are higher than the self-reported odds ratios for diabetes

#### Table 4

Regression results for establishing correction equations for body mass index (BMI), by sex, Full and Reduced Models generated from split-sample A, household population aged 18 years or older, 2005

	Variable	Coefficient	95% confidence interval
Men			
Full Model R <sup>2</sup> = 0.86 R <sup>2</sup> (adj) = 0.86	Intercept Self-reported BMI (kg/m <sup>2</sup> ) Aged 25 to 34 years <sup>†</sup> Aged 35 to 44 years <sup>†</sup> Aged 45 to 54 years <sup>†</sup> Aged 55 to 64 years <sup>†</sup> Aged 65 to 74 years <sup>†</sup> Aged 75 years or older <sup>†</sup> Dissatisfied with life Perceives self as underweight	-0.67 1.04* 0.64* 0.31 0.39 1.28* 1.16* 0.86* -0.97* -0.73*	-1.8 to 0.5 1.0 to 1.1 0.2 to 1.1 -0.2 to 0.8 -0.2 to 1.0 0.7 to 1.9 0.6 to 1.7 0.3 to 1.4 -1.6 to -0.3 -1.3 to -0.1
Reduced Model $R^2$ = 0.85 $R^2$ (adj) = 0.85	Intercept Self-reported BMI (kg/m <sup>2</sup> )	-1.08 1.08*	-2.2 to 0.0 1.0 to 1.1
Women Full Model R <sup>2</sup> = 0.92 R <sup>2</sup> (adj) = 0.92	Intercept Self-reported BMI (kg/m <sup>2</sup> ) Highest level of education is secondary graduation <sup>1</sup> Highest level of education is some postsecondary <sup>1</sup> Highest level of education is postsecondary graduation Perceives self as overweight End-digit preference (0 and 5)	1.01 1.01* -0.91* -0.32 0.70* 0.29*	-0.6 to 2.6 0.9 to 1.1 -1.5 to -0.3 -1.3 to 0.7 -1.0 to 0.0 0.2 to 1.2 0.0 to 0.6
$\begin{array}{l} \textbf{Reduced Model} \\ R^2 = 0.91 \\ R^2(adj) = 0.91 \end{array}$	Intercept Self-reported BMI (kg/m <sup>2</sup> )	-0.12 1.05*	-1.5 to 1.3 1.0 to 1.1

reference group is ages 18 to 24 years

<sup>‡</sup> reference group is less than secondary graduation

\* p<0.05

Note: Dependent variable is BMI based on measured height and weight.

Source: 2005 Canadian Community Health Survey.

in Models 1 and 2, and for high blood pressure, in Models 3 and 4.

Measured height and weight data were available for only a subsample of the 2005 CCHS. The ultimate goal of developing correction equations is to be able to apply them to the broader survey. When applied to the full sample of the 2005 CCHS (without different adjustments for telephone and in-person interviews) for respondents who were 18 years or older and who were not pregnant or breastfeeding (n=118,383), the models generated obesity estimates similar to, although slightly lower than, the measured values (Table 10). Based on data from both split-sample A and B, the self-reported prevalence of obesity was 16% for both sexes, while the measured prevalence was 25.6% for men and 22.3% for women. The models

generated obesity rates of approximately 23% for men and 21% for women.

#### Limitations

The response rate for the measured height and weight subsample of the CCHS was only 65%. If people who agreed to participate had different height and weight profiles than did those who refused, the sample could be biased. The self-reported prevalence of obesity among everyone who was selected to have their height and weight measured was 15.9% - 19.1% of non-respondents and 14% of respondents. However, when the special sampling weight was applied to those who underwent the physical measures, the prevalence of obesity based on self-reported data fell to 15.2%, comparable to that for the entire subsample.12

Bias in self-reported height may be due to inconsistent rounding between self-reported and measured data. When half-inches were reported, interviewers asked respondents to round up to the nearest inch, but for the measured values, height was recorded to the nearest 0.5 cm. Moreover, because interviewers recorded self-reported height only in metres, it was impossible to determine how many people reported in feet and inches and thereby assess the extent of this rounding bias.

For measured weight, it is not known if interviewers consistently asked respondents to empty their pockets and remove their footwear. And for selfreported weight, it is not known if respondents reported their weight with or without clothing, since interviewers told them to report their weight without clothing only if they asked.

Although interviewers were trained in the correct procedures for measuring height and weight, and the weigh scales and measuring tapes were calibrated, intra- and inter- interviewer reliability was not assessed.

BMI is commonly used as a measure of obesity on population surveys, but it has limitations: it cannot distinguish between muscle mass and fat, nor does it consider fat distribution.<sup>26</sup>

Finally, the models generated for this article were limited to the variables collected in the CCHS. It is possible that additional variables that were not part of the survey could be associated with the bias in weight, height or obesity.

#### Discussion

BMI calculated from self-reported height and weight underestimates obesity prevalence. This has implications for our understanding of the burden of obesity and the relationship between obesity and obesity-related health conditions. This study examined the feasibility of applying correction factors to self-reported estimates to determine if they could be adjusted to more closely approximate measured values.

Statistics Canada, Catalogue no. 82-003-XPE • Health Reports, Vol. 19, no. 3, September 2008 Feasibility of establishing correction factors to adjust self-reported estimates of obesity • Methodological Insights

#### Table 5 Correction equations to adjust self-reported estimates of weight, height and body mass index (BMI), by sex, Full and Reduced Models, household population aged 18 years or older, 2005

Sex and Model	Equation
Men	
Full Model 1	Weight <sub>(measured)</sub> = -0.30+1.01(weight <sub>selfreported</sub> )+0.54(age 25-34)+0.39(age 35-44)+0.50(age 45-54)+1.69(age 55-64)+0.83(age 65-74) +0.39(75 or older)+1.16(overweight)-1.52(underweight)
	Height <sub>(measured)</sub> = 12.17+0.93(height <sub>self-reported</sub> )-1.48(age 25-34)-0.43(age 35-44)-1.23(age 45-54)-2.44(age 55-64)-2.87(age 65-74)-2.84(75 or older) +2.22(life dissatisfaction)
Full Model 2	$BMI_{(measured)} = -0.67 + 1.04(BMI_{self-reported}) + 0.64(age 25-34) + 0.31(age 35-44) + 0.39(age 45-54) + 1.28(age 55-64) + 1.16(age 65-74) + 0.86(75 \text{ or older}) + 0.97(life dissatisfaction) - 0.73(underweight)$
Reduced Model 3	Weight <sub>(measured)</sub> = -2.19+1.05(weight <sub>self-reported</sub> ) Height <sub>(measured)</sub> = 7.70+0.95(height <sub>self-reported</sub> )
Reduced Model 4	BMI <sub>(measured)</sub> = -1.08+1.08(BMI <sub>self-reported</sub> )
Women	
Full Model 1	Weight <sub>(measured)</sub> = -1.25+1.04(weight <sub>self-reported</sub> )+1.25(overweight)+0.52(end-digit preference)
	Height <sub>(measured)</sub> = 14.85+0.91(height <sub>self-reported</sub> )-1.20(age 25-34)-0.87(age 35-44)-0.59(age 45-54)-1.34(age 55-64)-1.42(age 65-74)-3.79(75 or older) -0.32(ethnicity E/SE Asian)-0.73(ethnicity other)-0.66(activity limitation)
Full Model 2	BMI <sub>(measured)</sub> = 1.01+1.01(BMI <sub>self-reported</sub> )-0.91(secondary graduation)-0.32(some postsecondary)-0.53(postsecondary graduation)+0.70(overweight) +0.29(end-digit preference)
Reduced Model 3	Weight <sub>(measured)</sub> = -2.14+1.07(weight <sub>self-reported</sub> ) Height <sub>(measured)</sub> = 8.05+0.95(height <sub>self-reported</sub> )
Reduced Model 4	BMI <sub>(measured)</sub> = -0.12+1.05(BMI <sub>self-reported</sub> )
Source: 2005 Canadian	Community Health Survey.

#### Table 6 Mean weight, height and body mass index (BMI) for measured, self-reported and corrected data generated from splitsample B, by sex, household population aged 18 years or older, 2005

					Corre	ected	
	Sample size	Self-reported	Measured	Model 1 (Full) Height and weight	Model 2 (Full) BMI	Model 3 (Reduced) Height and weight	Model 4 (Reduced) BMI
Mean height (centimetre	es)						
Men	942	176.35*	175.21	175.42		175.44	
Women	1,087	162.28*	161.71	161.73		161.73	
Mean weight (kilograms	)						
Men	, 947	81.44*	83.24	83.26		83.27	
Women	1,080	64.47*	66.91	66.76		66.75	
Mean BMI (kg/m²)							
Men	949	26.12*	27.09	27.00	27.05	26.98	27.03
Women	1,080	24.55*	25.73	25.60	25.69	25.58*	25.68

not applicable

significantly different from measured estimate (p < 0.05)

Source: 2005 Canadian Community Health Survey.

In each of the four models tested, and in all analyses undertaken, the corrected estimates provided more accurate measures of overweight and obesity than did the self-reported values. However, this was not the case for the normal weight category. The sensitivity values for the normal weight population fell to as low as 84% in men (a 10percentage-point decrease) and to 83% in women (a 9-percentage-point decrease). Kuskowska-Wolk et al. also found a reduction in sensitivity for normal weight individuals.<sup>19</sup> We hypothesized that the decline in sensitivity was because heavier individuals have a greater reporting bias<sup>12</sup> (a greater tendency to underestimate their BMI), and therefore, different adjustments may be required depending on where the individual lies on the BMI distribution. Without these differing adjustments, sensitivity declines when a small proportion of

77

#### Table 7

Percentage distribution of population, by body mass index (BMI) category and sex, based on self-reported, measured and corrected data from split-sample B, household population aged 18 years or older, 2005

			Corrected						
BMI category	Self-reported	Measured	Model 1 (Full) Height and weight	Model 2 (Full) BMI	Model 3 (Reduced) Height and weight	Model 4 (Reduced) BMI			
Men									
Underweight	F	F	F	F	F	F			
Normal weight	43.1*	32.2	33.6	32.2	32.8	33.8			
Overweight	42.5	44.0	44.1	45.6	48.0	45.7			
Obese	13.8*	23.1	21.9	21.6	18.9*	20.1			
Women									
Underweight	4.7* <sup>E</sup>	3.1 <sup>E</sup>	2.7 <sup>E</sup>	1.5* <sup>E</sup>	2.7 <sup>E</sup>	1.9 <sup>E</sup>			
Normal weight	58.2*	46.9	46.8	47.0	46.6	47.1			
Overweight	24.6*	31.1	31.8	33.2	32.4	32.7			
Obese	12.5*	18.9	18.7	18.3	18.2	18.3			

interpret with caution (coefficient of variation between 16.6% and 33.3%)

significantly different from measured estimate (p < 0.05) F

Source: 2005 Canadian Community Health Survey.

#### Table 8

Sensitivity and specificity values for self-reported and corrected data, by sex, household population aged 18 years or older, 2005

	Underweight		No	Normal weight Overweight		verweight		Obese	Total	
	%	95% confidence interval	%	95% confidence interval	%	95% confidence interval	%	95% confidence interval	%	95% confidence interval
Sensitivity (% true positives) Men										
Self-reported Model 1 (Full - height and weight) Model 2 (Full - BMI) Model 3 (Reduced - height and weight) Model 4 (Reduced - BMI)		F F F F	93.9 87.8 85.5 83.8 85.8	91.7 to 96.2 83.6 to 91.4 80.8 to 90.1 77.6 to 90.1 79.7 to 91.9	71.1 79.8 81.1 82.8 81.1	66.2 to 76.0 73.2 to 86.4 74.5 to 87.7 76.5 to 89.0 74.7 to 87.6	58.7 76.0 74.6 70.2 73.8	51.7 to 65.7 67.1 to 84.9 65.5 to 83.8 59.9 to 80.4 64.8 to 82.8	75.0 81.2 80.7 79.9 80.7	72.0 to 78.0 77.3 to 85.1 76.7 to 84.7 75.6 to 84.1 76.6 to 84.7
Women Self-reported Model 1 (Full - height and weight) Model 2 (Full - BMI) Model 3 (Reduced - height and weight) Model 4 (Reduced - BMI)	77.8 66.8 <sup>e</sup> 39.3 <sup>e</sup> 66.8 <sup>e</sup> 45.6 <sup>e</sup>	63.2 to 92.3 42.6 to 91.2 18.3 to 60.4 43.2 to 90.3 24.0 to 67.5	91.8 85.1 83.4 85.6 86.9	88.9 to 94.8 79.0 to 91.3 77.2 to 89.6 79.6 to 91.7 81.4 to 92.4	62.6 74.3 75.0 77.1 79.7	56.8 to 68.5 67.7 to 81.0 68.4 to 81.5 70.7 to 83.6 73.9 to 85.5	68.5 86.1 85.1 85.4 86.0	62.3 to 74.8 78.7 to 93.5 77.5 to 92.7 78.0 to 92.8 78.6 to 93.4	77.8 81.4 79.7 82.4 83.2	74.9 to 80.7 77.4 to 85.3 75.7 to 83.8 78.5 to 86.2 79.6 to 86.8
Specificity (% true negatives)										
Men Self-reported Model 1 (Full - height and weight) Model 2 (Full - BMI) Model 3 (Reduced - height and weight) Model 4 (Reduced - BMI)	99.6 99.8 99.7 99.9 99.8	99.4 to 99.9 99.5 to 100.0 99.3 to 100.0 99.7 to 100.0 99.5 to 100.0	83.2 92.2 93.2 91.6 91.0	80.2 to 86.1 88.7 to 95.6 89.8 to 96.6 88.0 to 95.1 87.3 to 94.7	79.7 84.1 82.3 79.3 82.2	76.3 to 83.2 79.8 to 88.3 77.8 to 86.7 74.0 to 84.7 77.4 to 87.0	98.3 94.4 94.3 96.5 96.1	96.6 to 99.7 91.5 to 97.2 91.4 to 97.2 94.3 to 98.7 93.7 to 98.4	  	   
Women Self-reported Model 1 (Full - height and weight) Model 2 (Full - BMI) Model 3 (Reduced - height and weight) Model 4 (Reduced - BMI)	97.7 99.4 99.7 99.3 99.6	96.8 to 98.5 98.8 to 99.9 99.4 to 100.0 98.8 to 99.9 99.1 to 100.0	78.3 87.1 85.2 87.9 88.0	74.6 to 82.0 82.9 to 91.3 80.8 to 89.7 84.0 to 91.8 84.1 to 91.9	88.9 87.4 85.6 87.7 88.4	86.5 to 91.2 82.9 to 91.9 81.0 to 90.2 83.3 to 92.2 84.3 to 92.5	99.6 97.0 97.3 97.4 97.5	99.3 to 99.8 95.5 to 98.4 96.0 to 98.6 96.1 to 98.7 96.3 to 98.8	  	···· ···· ···

not applicable

E interpret with caution (coefficient of variation between 16.6% and 33.3%)

F too unreliable to be published (coefficient of variation greater than 33.3%)

Note: Reported estimates are based on data from split-samples A and B. Modelled estimates are generated from split-sample B.

Source: 2005 Canadian Community Health Survey.

#### Table 9

### Adjusted odds ratios relating self-reported, measured and corrected body mass index (BMI) to selected self-reported health conditions, household population aged 40 years or older, 2005

						Based on corrected values						
	Bas self-repo	sed on orted values	s measured values		Mod	el 1 (Full)	Model	Model 2 (Full)		(Reduced)	Model 4	(Reduced)
BMI category (range kg/m2)	Adjusted odds ratio	95% confidence interval										
Diabetes												
Normal weight (18.5 to 24.9)	1.0	 1 E to 4 0	1.0	 0 7 to 2 0	1.0	 0 0 to 2 2	1.0	 1 1 to 2 0	1.0	 1 1 to 2 0	1.0	 1 0 to 2 0
Overweight (25.0 to 29.9) Obese along $L(20.0 \text{ to } 24.9)$	2.0	1.5 l0 4.3	1.4	0.7 l0 2.8	1.8	0.9 l0 3.3	2.0	1.1 l0 3.8	1.ð 2.1.*	1.1 l0 3.0	2.0	1.2 l0 3.3
Obese class I (30.0 to 34.9) Obese class II and III (35.0 or more)	3.2 9.0*	4.5 to 17.9	2.2 5.9*	2.5 to 14.0	5.5 6.8*	3.7 to 12.5	3.9 7.3*	3.9 to 13.9	7.6*	4.0 to 14.2	3.2 7.4*	4.0 to 13.7
High blood pressure												
Normal weight (18.5 to 24.9)	1.0		1.0		1.0		1.0		1.0		1.0	
Overweight (25.0 to 29.9)	2.7*	1.9 to 3.8	2.1*	1.5 to 3.0	2.3*	1.6 to 3.2	2.5*	1.8 to 3.5	2.5*	1.7 to 3.5	2.4*	1.7 to 3.3
Obese class I (30.0 to 34.9)	4.2^	2.9 to 6.3	3.4° ⊑ 2*	2.3 to 5.2	4.0^	2.8 to 5.9	4.1^	2.8 to 6.0	4.5^	3.0 to 6.6	4./^ ⊑ ∠*	3.2 to 7.0
Obese class II and III (55.0 of 1101e)	0.0	3.2 (0 14.0	0.Z	2.9 10 9.3	0.0	5.5 10 10.7	0.0	5.4 10 10.5	0.1	5.4 10 10.9	0.0	3.2 10 9.0
Heart disease												
Normal weight (18.5 to 24.9)	1.0		1.0		1.0		1.0		1.0		1.0	
Overweight (25.0 to 29.9)	1.4	0.9 to 2.3	1.0	0.6 to 1.7	1.3	0.8 to 2.2	1.3	0.8 to 2.2	1.2	0.7 to 2.0	1.4	0.8 to 2.2
Obese class I (30.0 to 34.9)	1.6	1.0 to 2.6	1.5	0.8 to 2.9	1.2	0.7 to 2.0	1.4	0.8 to 2.4	1.3	0.8 to 2.2	1.5	0.9 to 2.5
Obese class II and III (35.0 or more)	3.7^	1.8 to 7.7	2.1	1.0 to 4.4	3.3^	1.8 to 6.2	3.4^	1.8 to 6.5	2.9^	1.5 to 5.6	2.8^	1.5 to 5.5
Arthritis												
Normal weight (18.5 to 24.9)	1.0		1.0		1.0		1.0		1.0		1.0	
Overweight (25.0 to 29.9)	1.2	0.8 to 1.7	1.2	0.8 to 1.7	1.5*	1.1 to 2.0	1.5*	1.1 to 2.1	1.5*	1.1 to 2.0	1.4*	1.0 to 1.9
Obese class I (30.0 to 34.9)	2.0*	1.3 to 3.0	1.2	0.8 to 1.8	1.7*	1.2 to 2.5	1.9*	1.3 to 2.8	1.9*	1.3 to 2.8	1.7*	1.2 to 2.5
Obese class II and III (35.0 or more)	3.1*	1.5 to 6.3	2.7*	1.6 to 4.6	3.5*	2.0 to 5.8	3.2*	1.8 to 5.4	3.2*	1.9 to 5.6	3.4*	1.9 to 6.0
Activity limitation												
Normal weight (18.5 to 24.9)	1.0		1.0		1.0		1.0		1.0		1.0	
Overweight (25.0 to 29.9)	1.2	0.9 to 1.6	1.2	0.9 to 1.6	1.2	0.9 to 1.7	1.1	0.8 to 1.5	1.0	0.8 to 1.4	1.1	0.8 to 1.5
Obese class I (30.0 to 34.9)	2.0*	1.3 to 3.0	1.5*	1.1 to 2.2	1.4	0.9 to 2.0	1.4	1.0 to 2.1	1.5*	1.0 to 2.1	1.5*	1.0 to 2.2
Obese class II and III (35.0 or more)	4.3*	2.2 to 8.2	2.9*	1.7 to 4.7	4.2*	2.6 to 6.8	3.7*	2.3 to 6.1	3.9*	2.4 to 6.5	3.1*	1.8 to 5.2
Fair/poor self-perceived health												
Normal weight (18.5 to 24.9)	1.0		1.0		1.0		1.0		1.0		1.0	
Overweight (25.0 to 29.9)	1.3	0.9 to 2.0	0.8	0.5 to 1.2	1.1	0.7 to 1.6	1.1	0.8 to 1.7	1.0	0.7 to 1.5	1.0	0.7 to 1.5
Ubese class I (30.0 to 34.9)	2.8*	1.8 to 4.3	1.7*	1.0 to 2.7	1.6*	1.0 to 2.5	1.7*	1.1 to 2.7	2.1*	1.3 to 3.3	2.1*	1.4 to 3.3
Obese class II and III (35.0 or More)	4.5	2.0 10 10.2	2.9"	1.0 10 5.2	4.1"	2.4 to 7.0	4.3	2.4 10 7.8	3.5	1.9 10 0.5	3.6	2.0 10 6.6

\* significantly different from estimate for normal weight category (p < 0.05)

. not applicable

Notes: Models control for age (continuous) and sex. Odds ratios for underweight group not reported because of small sample sizes.

Source: 2005 Canadian Community Health Survey.

normal weight individuals are erroneously shifted to the overweight category. We attempted to address this by incorporating polynomial regressions (quadratic terms for self-reported weight) and spline regression to determine if different slopes could be generated for different weight ranges. The quadratics and differential slopes were not significant, and we were unable to refine the estimates for those in the normal weight range. Therefore, although the adjustments improve the estimates for those who are overweight or obese, the non-adjusted numbers provide better estimates for respondents in the normal weight category because

the reporting bias is smaller in this group. Further research is needed to better understand how to improve selfreported overweight and obesity estimates without decreasing sensitivity for those in the normal weight range. More research is also required to determine if differential adjustments are necessary for respondents who were interviewed by telephone.

Despite this drawback, the improvement in classification for overweight and obese individuals is significant, and thus, we recommend the use of corrected estimates in addition to self-reported values in studies examining overweight and obesity in the adult population of the 2005 CCHS. We attempted to adjust for independent variables that were related to the reporting bias, but the R<sup>2</sup> of the Full Models (Models 1 and 2) was either the same as or only slightly higher than that of the Reduced Models (Models 3 and 4, which used only weight, height or BMI). In most cases, including the extra variables offered no predictive advantage. Plankey et al.<sup>31</sup> also found that more complex models (including self-reported BMI and additional covariates) only minimally improved predictive ability. Of the models we tested, all four generated similar means, prevalence rates and sensitivity values;

Table	10
-------	----

				Corre	ected	
BMI category	Self-reported	Measured	Model 1 (Full) Height and weight	Model 2 (Full) BMI	Model 3 (Reduced) Height and weight	Model 4 (Reduced) BMI
Males						
Underweight	0.7 <sup>E</sup>	0.9 <sup>E</sup>	1.0	1.2	0.9	1.0
Normal weight	41.8	32.4	31.2	30.0	31.2	33.1
Overweight	41.2	41.1	44.3	45.4	44.9	42.9
Obese	16.3	25.6	23.5	23.4	23.0	23.1
Females						
Underweight	4.3	2.6 <sup>E</sup>	2.7	2.1	2.5	2.0
Normal weight	54.0	46.1	46.5	46.4	46.8	46.7
Overweight	26.1	29.1	29.9	30.6	30.1	30.6
Obese	15.7	22.3	20.9	21.0	20.7	20.8

Percentage distribution of population, by body mass index (BMI) category and sex, when corrected estimates were applied to full 2005 Canadian Community Health Survey Sample, household population aged 18 years or older

<sup>E</sup> coefficient of variation between 16.6% and 33.3% (interpret with caution)

Notes: Measured and reported values were generated based on subsample of respondents whose height and weight were measured.

Source: 2005 Canadian Community Health Survey.

# What is already known on this subject?

- Self-reported data underestimate the true prevalence of obesity and overestimate the relationship between obesity and obesityrelated health conditions.
- For fiscal and logistical reasons, most population health surveillance in Canada is based on self-reported information.

## What does this study add?

- Correction factors can be generated to adjust self-reported data to produce more accurate estimates of obesity.
- Although not perfectly predictive of the measured values of body mass index (BMI), corrected values are an improvement over self-reported estimates.
- For future studies examining BMI for adult populations, based on data from the 2005 Canadian Community Health Survey, corrected estimates of BMI are recommened.

no model stood out as being consistently superior. Model 4, however, had the further advantage of being the most parsimonious, and therefore, showing the greatest utility if it is determined that the equations are generalizable.

This method of generating corrected estimates (linear regression with measured BMI as the outcome) has been used in the past, 10, 19, 31-34 but to our knowledge, has never been attempted on data for the Canadian population. Plankey et al.<sup>31</sup> concluded that a systematic error was associated with the reporting bias, which was impossible to correct with this method. However, in their work, the self-reported sensitivity values for the obese population (BMI 27.3 kg/m<sup>2</sup> or more) were 80% in men and 85% in women and increased only marginally with the corrected models. By contrast, in the current study, self-reported sensitivity of obesity was much lower-59% for men and 69% for womenand the correction equations increased these values significantly. Also, the reporting bias in our study was two to three times larger than that in the 1976-1980 NHANES II, on which the analysis of Plankey et al.was based.

The generalizability of these equations has not been determined. Some authors<sup>33</sup> assume transportability, while others<sup>30</sup> have shown that correction equations are applicable only to the population for which they have been established. In one Swedish study,<sup>32</sup> researchers demonstrated that because height was under- rather than over-reported in that country, selfreported estimates of BMI did not require calibration.

More research using Canadian data is required to determine if these equations are stable across Canadian populations and over time. It is probable that the increase in obesity in recent years<sup>35</sup> has been accompanied by a corresponding increase in reporting bias, which could indicate temporal instability in the equations. At least one study that has examined the bias over time has found that it has increased.<sup>36</sup>

In the interim, surveys that collect self-reported and measured height and weight would benefit from standardization of protocols to ensure that equipment is regularly calibrated and that respondents are asked to report their weight in a consistent way and are measured in light clothing, without shoes. Rounding should also be minimized, if not eliminated.

#### Conclusion

Although measured data for height and weight provide the most accurate estimates of the prevalence of obesity based on BMI, the costs of collecting such data are often prohibitive for large population-based surveys. Corrected estimates, though not identical to measured BMI values, are a significant improvement over estimates based on self-reported data, which substantially underestimate obesity prevalence and overestimate the relationship between obesity and disease.

#### Acknowledgements

The authors thank Julie Bernier for her methodological assistance and members of the Health Information and Research Division at Statistics Canada for their input on this research.

### References

- World Health Organization. Obesity and Overweight. Geneva: World Health Organization; 2006 [cited 2008 May 7]. Available from: http://www.who.int/ mediacentre/factsheets/fs311/en/ index.html.
- 2. Tjepkema M. Adult obesity. *Health Reports* (Statistics Canada, Catalogue 82-003) 2006; 17(3): 9-25.
- 3. Shields M. Overweight and obesity among children and youth. *Health Reports* (Statistics Canada, Catalogue 82-003) 2006; 17(3): 27-42.
- Rigby N, Leach R., James WPT. Seeking bold solutions for Britain's runaway obesity epidemic. International Obesity Taskforce Briefing Paper. London: IOTF, 2003.
- 5. Sturm R, Ringel JS, Andreyeva T. Increasing obesity rates and disability trends. *Health Affairs* 2004; 23(2): 199-205.
- 6. Katzmarzyk PT, Janssen I. The economic costs associated with physical inactivity and obesity in Canada: an update. *Canadian Journal of Applied Physiology* 2004; 29: 90–115.
- Engstrom JL, Paterson SA, Doherty A, et al. Accuracy of self-reported height and weight in women: an integrative review of the literature. *Journal of Midwifery and Women's Health* 2003; 48: 338–45.
- 8. Niedhammer I, Bugel I, Bonenfant S, et al. Validity of self-reported weight and height in the French GAZEL cohort. *International Journal of Obesity and Related Metabolic Disorders* 2000; 24: 1111–18.

- Nieto-Garcia FJ, Bush TL, Keyl PM. Body mass definitions of obesity: sensitivity and specificity using selfreported weight and height. *Epidemiology* 1990; 1: 146–52.
- Rowland ML. Self-reported weight and height. *The American Journal of Clinical Nutrition* 1990; 52: 1125–33.
- 11. Connor Gorber S, Tremblay M, Moher D, Gorber B. A comparison of direct vs. self-report measures for assessing height, weight and body mass index: a systematic review. *Obesity Reviews* 2007; 8: 307-26.
- Shields M, Connor Gorber S, Tremblay M. Estimates of obesity based on selfreport versus direct measures. *Health Reports* (Statistics Canada, Catalogue 82-003) 2008; 19(2): 61-76.
- 13. Puhl RM, Brownell KD. Psychosocial origins of obesity stigma: toward changing a powerful and pervasive bias. *Obesity Reviews* 2003; 4: 213-27.
- Kumanyika S, Jeffery RW, Morabia A, et al. Obesity prevention: the case for action. *International Journal of Obesity* 2002; 26: 425-36.
- 15. Canadian Institute for Health Information. Improving the Health of Canadians: Promoting Healthy Weights. Ottawa: Canadian Institute for Health Information, 2006.
- Shields M, Connor Gorber S, Tremblay M. Effects of measurement on obesity and morbidity. *Health Reports* (Statistics Canada, Catalogue 82-003) 2008; 19(2): 77-84.

- Yannakoulia M, Panagiotakos DB, Pitsavos C, Stefanadis C. Correlates of BMI misreporting among apparently healthy individuals: the ATTICA study. *Obesity* 2006; 14(5): 894-901.
- Chiolero A, Peytremann-Bridevaux I, Paccaud F. Associations between obesity and health conditions may be overestimated if self-reported body mass index is used. *Obesity Reviews* 2007; 8: 373-4.
- 19. Kuskowska-Wolk A, Bergstrom R, Bostrom G. Relationship between questionnaire data and medical records of height, weight and body mass index. *International Journal of Obesity* 1992; 16: 1-9.
- Statistics Canada. Population health surveys. Ottawa: Statistics Canada; 2005 [cited 2008 Jan 20]. Available from: http://www.statcan.ca/english/ concepts/hs/index.htm#content.
- Béland Y. Canadian Community Health Survey – methodological overview. *Health Reports* (Statistics Canada, Catalogue 82-003) 2002; 13(3): 9-14.
- 22. Ziebland S, Thorogood M, Fuller A, Muir J. Desire for the body normal: body image and discrepancies between self-reported and measured height and weight in a British population. *Journal* of Epidemiology and Community Health 1996; 50: 105–6.
- 23. Bostrom G, Diderichsen F. Socioeconomic differentials in misclassification of height, weight and body mass index based on questionnaire data. International Journal of Epidemiology 1997; 26: 860-6.

- 24. Roberts RJ. Can self-reported data accurately describe the prevalence of overweight? *Public Health* 1995;109: 275-84.
- 25. World Health Organization. *Physical Status: The Use of and Interpretation of Anthropometry, Report of the WHO Expert Committee* (WHO Technical Report Series, No. 854) Geneva: World Health Organization, 1995.
- 26. Health Canada. *Canadian Guidelines* for Body Weight Classification in Adults (Catalogue H49-179) Ottawa: Health Canada, 2003.
- Rao JNK, Wu CFJ, Yue K. Some recent work on resampling methods for complex surveys. *Survey Methodology* (Statistics Canada, Catalogue 12-001) 1992; 18(2): 209-17.
- 28. Rust KF, Rao JNK. Variance estimation for complex surveys using replication techniques. *Statistical Methods in Medical Research* 1996; 5: 281-310.

- 29. Statistics Canada. Canadian Community Health Survey Cycle 3.1 Derived Variable Specifications. Ottawa: Statistics Canada, 2006.
- 30. Visscher TLS, Viet AL, Kroesbergen HT, Seidell JC. Underreporting of BMI in adults and its effect on obesity prevalence estimations in the period 1998 to 2001. *Obesity* 2006; 14(11): 2054-63.
- Plankey MW, Stevens J, Flegal KM, Rust PF. Prediction equations do not eliminate systematic error in selfreported body mass index. *Obesity Research* 1997; 5(4): 308-14.
- 32. Bolton-Smith C, Woodward M, Tunstall-Pedoe H, Morrison C. Accuracy of the estimated prevalence of obesity from self reported height and weight in an adult Scottish population. *Journal of Epidemiology and Community Health* 2000; 54:143-8.

- 33. Cawley J. The impact of obesity on wages. *The Journal of Human Resources* 2004; 39(2): 451-74.
- Nyholm M, Gullberg B, Merlo J, et al. The validity of obesity based on selfreported weight and height: implications for population studies. *Obesity* 2007; 15(1): 197-208.
- 35. Katzmarzyk PT. The Canadian obesity epidemic, 1985–1998. *Canadian Medical Association Journal* 2002; 166: 1039-40.
- 36. Ezzati M, Martin H, Skjold S, et al. Trends in national and state-level obesity in the USA after correction for self-report bias: analysis of health surveys. *Journal of the Royal Society of Medicine* 2006; 99: 250-7.